Relationships between headache frequency, disability, and disability-related unemployment among adults with migraine

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Plain language summary

We used a large survey to study the relationship between how many headaches a person has and their job status. People with more than 15 headache days each month were almost 3 times more likely to say they were not working because of disability than people with less than 5 headache days each month. This is important for helping those living with migraine to have better quality of life and understand how migraine affects employment.

Implications for managed care pharmacy

These findings demonstrate a link between frequency of migraine attacks and unemployment. This is important for employers and policymakers. Our findings can inform the design of targeted interventions for migraine prevention to facilitate the gainful employment of adults with migraine and to help guide decisions on disability benefits eligibility.

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ABSTRACT

BACKGROUND: Migraine is the second most common cause of disability worldwide. Understanding the relationship between migraine and employment status is critical for policymakers, as disability-related unemployment is associated with eligibility for private or governmental disability insurance payments and other associated support for those unable to work because of disability.

OBJECTIVE: To assess the association between migraine frequency and selfreported employment status and overall disability in a US representative survey.

METHODS: Using data from the 2019 National Health and Wellness Survey (NHWS) (Kantar Health), adults in the United States (aged 18-65 years) reporting at least 1 migraine day in the past 30 days were categorized by headache frequency: low-frequency episodic migraine (LFEM) (≤4 days/month), moderate-frequency EM (MFEM) (5-9 days/month), high-frequency EM (HFEM) (10-14 days/month), or chronic migraine (CM) (≥15 days/month). A control group of adults without migraine with similar baseline characteristics was identified by propensity score matching. Disabilityrelated unemployment was defined as participants responding "short-term disability" or "long-term disability" to occupational status on the NHWS. The frequency of shortor long-term disability was then evaluated across headache frequency groups. In addition, participants were asked to assess migraine-related disability via the Migraine Disability questionnaire (MIDAS).

RESULTS: A total of 1,962 respondents with LFEM, 987 with MFEM, 554 with HFEM, and 926 with CM were included in this

analysis, along with 4,429 matched controls. Headache frequency was associated both with increased MIDAS score and with employment disability (P < 0.001); 12.3% (n=114 of 926) of participants with CM reported employment disability, as did 4.4% (n=86 of 1,962) of the LFEM group and 6.9% (n=306 of 4,429) of matched controls. There was considerable discordance between the proportion of participants classified as disabled via MIDAS vs those reporting employment-related disability.

CONCLUSIONS: More frequent migraine headaches are associated with a higher likelihood of self-reported short- and long-term employment disability and overall migrainerelated disability, suggesting that health and economic policymakers must seek ways to maximize the employment opportunities for people living with migraine that may benefit from novel preventive treatments. Disability is defined by the International Classification of Functioning, Disability, and Health as impairments or limitations to activity participation caused by a health condition that alter an individual's relationship with themselves and/or their personal/working environment.1-4 Medical and health services researchers typically assess disability by quantifying impairment associated with a disease. For example, migraine is the second-ranked cause of disability worldwide and the most common cause of years lived with disability (YLD) among those aged 15-49,5 with YLD generally increasing with greater disease severity.6 In particular, migraine contributes to reduced overall labor force participation, increased long-term disability, and overall work impairment.^{7,8} In addition, according to data from the Baltimore County Migraine Study, MEDSTAT's MarketScan, medical claims, and statistics from the Census Bureau and the Bureau of Labor Statistics, patients with migraine reported approximately 112 million bedridden days (ie missed or impaired work days) per year.9

Migraine-associated disability is frequently measured using validated patient-reported outcomes, such as the Migraine Disability Assessment (MIDAS) questionnaire.¹⁰⁻¹² MIDAS consists of 5 questions that measure the extent of disability across 3 contextual domains: school or work for pay; household role; and participation in social, family, and leisure activities. The responses to each question are provided in number of days, which are then totaled to determine the level of disability: "little or no disability" (0-5 days); "mild disability" (6-10 days); "moderate disability" (11-20 days); and "severe disability" (21-40 days).^{11,13,14}

Although the extent of disability in migraine has been characterized across multiple domains, including occupational, academic, familial, and social,¹⁵⁻¹⁸ and the disabling effects of migraine are known to contribute to large direct and indirect economic costs,^{9,19-20} this body of research has not fully elucidated the relationship between disability and the impact of migraine on employment status (ie, disability-related unemployment).²¹⁻²³ Understanding this relationship is critical for policymakers, as unemployment that is a function of impairment may imply eligibility for private or governmental disability insurance payments and other associated support for those physically or mentally unable to work.

In the current study, we aimed to explore the relationship between headache frequency and disability-related unemployment among a large population-based cohort of adults living with migraine in the United States. We also examined the association between disability and employment status.

Methods

STUDY DESIGN

This retrospective cross-sectional study included data from the 2019 National Health and Wellness Survey (NHWS) from respondents in the United States. The NHWS is a selfadministered, Internet-based survey conducted annually (Cerner Enviza, North Kansas City, MO) in the United States and several other countries.²⁴⁻²⁷ The NHWS collects information on more than 200 health conditions, including a module dedicated to migraine, which assesses migraine-related symptomology, interactions with health care providers (eg, physicians, physician assistants, and nurse practitioners), and medication use.28 All respondents, regardless of migraine status, provided information on demographic and baseline health factors. Survey respondents were recruited through participation in opt-in online survey panels, with quota sampling within the survey panel to ensure countryspecific representativeness in age, sex, and race and ethnicity distributions based on the US Census.

The NHWS was granted exemption by the Pearl Institutional Review Board (Indianapolis, IN) because of the low risk to participants of its deidentified/anonymized secondary data. The current study was conducted in accordance with the Good Pharmacoepidemiology Practices guidelines issued by the International Society for Pharmacoepidemiology.²⁹

STUDY SAMPLE

From the total 2019 NHWS survey (N=74,994), respondents aged 18-65 years with a self-reported physician diagnosis of migraine who reported experiencing at least 1 migraine attack in the past 30 days (n=4,487, total eligible respondents) were analyzed in aggregate and stratified by headache frequency (<u>Supplementary Figure 1</u>, available in online article). Episodic migraine (EM) and chronic migraine (CM) were defined using the cutoff of 15 headache days per month, as recommended by the International Classification of Headache Disorders, Third Edition (ICHD-3).³⁰

To facilitate comparisons to individuals without migraine or nonmigraine controls, we selected an additional cohort of adults from the NHWS who had not been diagnosed with migraine and did not report experiencing migraine symptoms in the past year (n=45,962 total eligible respondents). Thus, the impact of headache frequency could be characterized within the migraine cohort, in addition to being contextualized against the wider nonmigraine group.

MEASURES

Migraine Frequency. EM subtypes were defined in accordance with ICHD-3 criteria. Low-frequency EM (LFEM) (n=1,971 total eligible respondents): no more than 4 headache days/month, moderate-frequency EM (MFEM) (n=998): 5-9 headache days/month, and high-frequency EM (HFEM) (n=565): 10-14 headache days/month.^{31,32} Nevertheless, the definition used for this study does not include "migraine days" because the information necessary to characterize a headache episode due to migraine is not available in the NHWS, as the reported headaches could have been caused by reasons other than migraine.

Disability Status. Disability status was characterized based on 2 different measures: (1) self-reported employment status and (2) MIDAS score.

Self-reported employment status. In the NHWS, respondents were asked, "What is your employment status?" Choices were employed full-time, self-employed, employed part-time, homemaker, retired, student, short-term or longterm disability, not employed but looking for work, or not employed and not looking for work. Respondents selected the single response that best characterized their current employment. We assumed that those who selected shortterm or long-term disability self-assessed that they were unable to work for some period of time because of impairments (migraine or other existing comorbidities) and might therefore be eligible for entitlements such as sick leave, Social Security benefits, workers' compensation, or private disability insurance. All survey participants in both the matched control population and migraine population reported on employment status.

MIDAS score. For the purposes of this study, individuals with higher MIDAS scores were defined as having greater levels of disability as a result of disruptive headache days.^{11,13,14} These scores were further categorized as indicating no disability (0-5), mild disability (6-10), moderate disability (11-20), or severe disability (\geq 21).¹³ We assumed that, in this context, "disability" as measured by the MIDAS may be a more general assessment of disability beyond employment status. Of the respondents with migraine, 100% completed the MIDAS portion of the survey.

Demographics. Demographic characteristics reported from the NHWS included age, sex, employment status, race and ethnicity, marital status, education, household income, and insurance status. Self-reported clinical characteristics of the migraine sample included years since diagnosis, monthly headache days, current use of prescription and over-the-counter medications, duration of medication use, and prescribing patterns (eg, types of prescribers, reasons for adding/switching medications).

General Health and Lifestyle Characteristics. General health variables reported from the NHWS included body

mass index, smoking status, alcohol use, and exercise behavior. The Charlson Comorbidity Index (CCI) was used to capture overall comorbidity burden.³³ Given the high prevalence of migraine among women, we also examined the comorbidity burden associated with key women's health conditions available within the NHWS. The total number of diagnoses per person (range 0-11) were tallied across the following conditions: breast cancer, cervical cancer, ovarian cancer, uterine cancer, endometriosis, fibroids, dysmenorrhea, heavy menstrual bleeding, hot flashes, premenstrual dysphoric disorder, and premenstrual syndrome. The use of oral contraceptives was not assessed.

Mental Health Characteristics. Anxiety and depression are well known to be comorbid with migraine and are known to play a role in transformation of migraine from episodic to chronic.³⁴⁻³⁷ We explored both anxiety and depression in 2 ways: via self-reported diagnoses available in the NHWS and via symptom severity, as assessed by the Generalized Anxiety Disorder questionnaire (7-item instrument) (GAD-7)38 and the Patient Health Questionnaire (9-item instrument) (PHQ-9).39 Among diagnosed patients, prevalence of anxiety/depression was subdivided by medication usage: patients who reported currently using a medication to treat their anxiety/depression were coded as "treated," whereas patients who did not report using any anxiety/ depression medication were coded as "untreated." For all patients, irrespective of self-reported diagnosis, symptom severity was ascertained via the GAD-7 and the PHQ-9. Both instruments query about anxiety/depressive symptoms experienced in the past 2 weeks, with items scored on a 0 (not at all) to 3 (nearly every day) scale. Scores of 5, 10, 15, and 20 serve as cutoffs indicating mild, moderate, moderately severe, and severe levels of depression on the PHQ-9. Scores of 5, 10, and 15 serve as cutoffs indicating mild, moderate, and severe anxiety on the GAD-7.

STATISTICAL ANALYSIS

Unadjusted Bivariate Analyses. Propensity score matching was used to generate comparable cohorts of adults with and without migraine that possessed similar demographic and general health characteristics. First, bivariate analyses were performed to identify baseline characteristics (shown in Table 1) that were unbalanced between the migraine and matched control cohorts. This analysis yielded 14 variables that were included as covariates in the propensity score model: age (mean), sex (male, female), race and ethnicity (non-Hispanic White, non-Hispanic Black, Hispanic, other), marital status (married/living with partner, other), university education (4-year degree, <4-year degree), household income (<\$75k, >\$75k), insurance status (insured, uninsured), CCI (mean), body mass index (overweight/obese, TABLE 1

Demographics and Health Characteristics

| | Migraine frequency | | | | | | | | | | | | |
|--------------------------------|--------------------------------|--------------------|-------------------------------|--------------------|---------------------------------|--------------------|----------------------------------|------------------|---|----------------------|----------------------|------------|------------|
| | LFEM (≤4 days) (n=1,962) | | MFEM (5-9 days) (n=987) | | HFEM (10-14 days) (n=554) | | Chronic (≥15 days) (n=926) | | Matched control (no migraine) (n=4,429) | | Total (N = 8,858) | | |
| | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | P value |
| Age (years) | 41.2 _a | 12.9 | 39.5 _b | 12.8 | 39.6 _{a,b} | 12.8 | 41.2 _a | 12.9 | 40.9 _a | 13.8 | 40.8 | 13.4 | 0.002 |
| Sex | | | | | | | | | | | | | |
| Male | 26.9% | 528 _a | 23.8% | 235 _{a,b} | 23.1% | 128 _{a,b} | 21.0% | 194 _b | 25.7% | 1,140 _a | 25.1% | 2,225 | 0.005 |
| Female | 73.1% | 1,434 _a | 76.2% | 752 _{a,b} | 76.9% | 426 _{a,b} | 79.1% | 732 _b | 74.3% | 3,289 _a | 74.9% | 6,633 | 0.005 |
| Race and ethnicity | | | | | | | | | | | | | |
| African American | 10.1% | 198 _a | 9.0% | 89 _{a,b} | 9.2% | 51 _{a,b} | 6.5% | 60 _b | 9.9% | 437 _a | 9.4% | 835 | |
| American Indian | 1.0% | 20 _a | 0.7% | 7 _a | 0.9% | 5 _a | 1.3% | 12 _a | 0.6% | 25 _a | 0.8% | 69 | |
| Asian | 4.8% | 95 _a | 4.1% | 40 _a | 3.1% | 17 _{a,b} | 1.6% | 15 _b | 5.1% | 228 _a | 4.5% | 395 | |
| Hispanic | 13.4% | 263 _a | 12.8% | 126 _a | 15.0% | 83 _a | 13.2% | 122 _a | 15.1% | 668 _a | 14.2% | 1,262 | <0.001 |
| Non-Hispanic White | 66.4% | 1,303 _a | 69.8% | 689 _{a,b} | 68.2% | 378 _{a,b} | 72.5% | 671 _b | 65.4% | 2,896 _a | 67.0% | 5,937 | |
| Mixed | 3.0% | 59 _a | 2.8% | 28 _a | 2.9% | 16 _a | 3.5% | 32 _a | 3.1% | 137 _a | 3.1% | 272 | |
| Other | 1.2% | 24 _a | 0.8% | 8 _a | 0.7% | 4 _a | 1.5% | 14 _a | 0.9% | 38 _a | 1.0% | 88 | |
| Marital status | | | | | | | | | | | | | |
| Single/not living with partner | 42.0% | 825 _a | 44.8% | 442 _a | 42.1% | 233 _a | 46.4% | 430 _a | 45.5% | 2,014 _a | 44.5% | 3,944 | |
| Married/living with partner | 57.8% | 1,135 _a | 55.2% | 545 _a | 57.8% | 320 _a | 53.3% | 494 _a | 54.2% | 2,400 _a | 55.2% | 4,894 | 0.05 |
| Decline to answer | 0.1% | 2 _a | 0% | 0ª | 0.2% | 1 _a | 0.2% | 2 _a | 0.3% | 15 _a | 0.2% | 20 | |
| Education | | | | | | | | | | | | | |
| Less than university education | 36.4% | 714 _a | 39.4% | 389 _{a,c} | 41.9% | 232 _{a,c} | 51.3% | 475 _b | 42.8% | 1,897 _c | 41.8% | 3,707 | |
| University education or higher | 51.1% | 1,002 _a | 47.9% | 473 _{a,b} | 42.8% | 237 _b | 35.0% | 324 _c | 43.8% | 1,942 _{b,d} | 44.9% | 3,978 | <0.001 |
| Decline to answer | 12.5% | 246 _a | 12.7% | 125 _a | 15.3% | 85 _a | 13.7% | 127 _a | 13.3% | 590 _a | 13.2% | 1,173 | 1 |
| Annual household income | | | | | | | | | | | | | |
| <\$25K | 15.7% | 308 _a | 17.9% | 177 _{a,c} | 17.9% | 99 _{a,c} | 26.8% | 248 _b | 19.1% | 848 _c | 19.0% | 1,680 | |
| \$25K to <\$50K | 22.0% | 432 _a | 25.2% | 249 _{a,b} | 27.1% | 150 _{a,b} | 27.9% | 258 _b | 24.9% | 1,105 _{a,b} | 24.8% | 2,194 | |
| \$50K to <\$75K | 19.5% | 382 _a | 16.7% | 165 _a | 19.1% | 106 _a | 17.0% | 157 _a | 18.6% | 825 _a | 18.5% | 1,635 | <0.001 |
| ≥\$75K | 40.0% | 784 _a | 37.2% | 367 _{a,c} | 33.6% | 186 _{a,c} | 25.2% | 233 _b | 33.7% | 1,493 _c | 34.6% | 3,063 | |
| Decline to answer | 2.9% | 56, | 2.9% | 29, | 2.3% | 13, | 3.2% | 30 _a | 3.6% | 158, | 3.2% | 286 | 1 |

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not overweight/obese), smoking status (current/former smoker, nonsmoker), exercise status (exercise, do not exercise), comorbid anxiety (treated, untreated, none), comorbid depression (treated, untreated, none), and region of residence (northwest, midwest, south, west). This list represents the variables that exhibited the largest differences across groups and/or the variables of most a priori relevance in relation to disability (eg, CCI as a measure of comorbidity burden, which could additively contribute to disability aside from headache frequency).

Matched Bivariate Analyses. Using recommended methods,⁴⁰ the above covariates were entered into a logistic regression model predicting the presence/absence of migraine. Respondents' regression-estimated probabilities were saved and used as propensity scores. Following a 1:1

| TABLE 1 Demogra | phics a | nd Hea | alth Cha | racteri | i stics (c | ontinue | ed) | | | | | | |
|--|--------------------------------|--------------------|-------------------------------|--------------------|---------------------------------|--------------------|----------------------------------|--------------------|---|----------------------|----------------------|------------|------------|
| | | | | | Μ | ligraine | frequenc | у | | | | | |
| | LFEM (≤4 days) (n=1,962) | | MFEM (5-9 days) (n=987) | | HFEM (10-14 days) (n=554) | | Chronic (≥15 days) (n=926) | | Matched control (no migraine) (n=4,429) | | Total (N = 8,858) | | |
| | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | P value |
| Currently employed (FT/PT/SE) | 71.5% | 1,402 _a | 69.9% | 690 _a | 67.0% | 371 _{a,c} | 54.1% | 501 _b | 64.0% | 2,834 _c | 65.5% | 5,798 | < 0.001 |
| Currently insured | 91.9% | 1,803 _a | 90.6% | 894 _a | 91.7% | 508 _a | 88.9% | 823 _a | 90.4% | 4,003 _a | 90.7% | 8,031 | 0.087 |
| Charlson Comorbidity Index (1987) | 0.5 _a | 1.0 | 0.6 _{a,b} | 1.1 | 0.6 _{b,c} | 1.2 | 0.7 _c | 1.2 | 0.6 _{b,c,d} | 1.3 | 0.6 | 1.2 | <0.001 |
| Women's health comorbidities ^b | 0.7 _a | 1.0 | 0.7 _{a,b} | 1.1 | 0.8 _b | 1.1 | 0.9 _{b,c} | 1.2 | 0.4 _d | 0.8 | 0.6 | 1.0 | <0.001 |
| BMI (kg/m²) | 28.1 _a | 7.8 | 28.4 _{a,b,d} | 7.8 | 29.2 _{b,c,d} | 8.2 | 30.0 _c | 8.5 | 28.8 _d | 7.9 | 28.8 | 8.0 | < 0.001 |
| Uses alcohol | 74.3% | 1,458 _a | 73.2% | 722 _a | 71.7% | 397 _{a,c} | 60.9% | 564 _b | 67.8% | 3,002 _c | 69.3% | 6,143 | < 0.001 |
| Uses tobacco | | | | | | | | | | | | | |
| Never smoked | 58.9% | 1,155 _a | 58.8% | 580 _a | 55.6% | 308 _a | 54.0% | 500 _a | 55.2% | 2,444 _a | 56.3% | 4,987 | _ |
| Former smoker | 20.0% | 393 _a | 18.7% | 185 _a | 21.7% | $120_{a,b}$ | 22.4% | 207 _{a,b} | 24.3% | 1,076 _b | 22.4% | 1,981 | <0.001 |
| Current smoker | 21.1% | 414 _a | 22.5% | 222 _a | 22.7% | 126 _a | 23.7% | 219 _a | 20.5% | 909 _a | 21.3% | 1,890 | |
| Engages in exercise | 73.8% | 1,447 _a | 73.5% | 725 _a | 71.1% | 394 _{a,b} | 64.5% | 597 _b | 68.1% | 3,015 _{b,c} | 69.7% | 6,178 | <0.001 |
| Comorbid anxiety | | | | | | | | | | | | | |
| Treated (Rx use) | 19.8% | 389 _a | 24.6% | 243 _b | 28.0% | 155 _{b,c} | 32.8% | 304 _c | 24.2% | 1,070 _b | 24.4% | 2,161 | |
| Untreated (no Rx use) | 15.3% | 300 _a | 17.8% | 176 _{a,b} | 17.0% | 94 _{a,b} | 21.2% | 196 _b | 18.1% | 802 _{a,b} | 17.7% | 1,568 | <0.001 |
| No anxiety | 64.9% | 1,273 _a | 57.6% | 568 _b | 55.1% | 305 _b | 46.0% | 426 _c | 57.7% | 2,557 _b | 57.9% | 5,129 | |
| Comorbid depression | | | | | | | | | | | | | |
| Treated (Rx use) | 22.5% | 442 _a | 26.3% | 260 _{a,c} | 26.0% | 144 _{a,c} | 35.2% | 326 _b | 27.1% | 1,199 _c | 26.8% | 2,371 | |
| Untreated (no Rx use) | 16.6% | 326 _a | 19.7% | 194 _{a,b} | 22.2% | 123 _b | 23.0% | 213 _{b,c} | 20.3% | 898 _{b,d} | 19.8% | 1,754 | <0.001 |
| No depression | 60.9% | 1,194 _a | 54.0% | 533 _b | 51.8% | 287 _b | 41.8% | 387 _c | 52.7% | 2,332 _b | 53.4% | 4,733 | |
| PHQ-9 score (0-27) | 7.4 _a | 7.0 | 8.7 _b | 6.9 | 10.1 _c | 7.4 | 11.0 _c | 7.6 | 7.2 _a | 7.0 | 8.0 | 7.2 | <0.001 |

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matching ratio, nearest neighbor matches were obtained using a 0.2 caliper restriction. To ensure that the matched control sample would be broadly representative of each of the migraine subgroups, 2 sets of matches were conducted: one to identify individuals from the matched control group similar to the LFEM plus MFEM migraine cohorts (n=2,949; <10 headache days per month) and a second to identify individuals from the matched control group similar to the HFEM plus CM cohorts (n=1,480; ≥10 headache days per month). These subgroups were combined into a common matched control cohort (n=4,429) for all subsequent analyses.

Standardized mean differences (SMDs) confirmed that the matched control cohort was acceptably matched to the aggregate migraine cohort (all SMDs <0.10),⁴¹ as well

as to each of the migraine frequency subgroups (all SMDs <0.25).^{42,43} See <u>Supplementary Table 1</u> and <u>Supplementary Table 2</u> in the Supplementary Materials for further details. Analyses were performed using IBM SPSS version 25.0, SAS version 9.4, and R version 3.6.0.

Descriptive analyses were conducted. Categorical variables were reported as counts and percentages and analyzed using chi-square tests; continuous variables were reported as means±SDs and analyzed using 1-way analysis of variance tests. P values were provided for the omnibus test, and statistically significant pairwise differences between the groups were denoted using subscripts. Two-tailed P values of <0.05 were considered statistically significant.

| TABLE 1 |
|---------|
|---------|

Demographics and Health Characteristics (continued)

| | Migraine frequency | | | | | | | | | | | | |
|---|--------------------------------|--------------------|---------------------------------|--------------------|---------------------------------|--------------------|----------------------------------|--------------------|---|--------------------|----------------------|------------|------------|
| | LFEM (≤4 days) (n=1,962) | | MFEM (5-9 days) (n = 987) | | HFEM (10-14 days) (n=554) | | Chronic (≥15 days) (n=926) | | Matched control (no migraine) (n=4,429) | | Total (N = 8,858) | | |
| | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | P value |
| PHQ-9 severity | | | | | | | | | | | | | |
| None or minimal (0-4) | 44.0% | 864 _a | 35.3% | 348 _b | 28.9% | 160 _{b,c} | 23.8% | 220 _c | 45.8% | 2,028 _a | 40.9% | 3,620 | - <0.001 |
| Mild depression (5-9) | 24.4% | 479 _a | 25.4% | 251 _a | 24.5% | 136 _a | 23.8% | 220 _a | 23.2% | 1,026 _a | 23.8% | 2,112 | |
| Moderate depression (10-14) | 14.4% | 282 _a | 18.5% | 183 _b | 18.6% | 103 _{a,b} | 21.6% | 200 _b | 14.6% | 647 _a | 16.0% | 1,415 | |
| Moderately severe depression (15-19) | 9.6% | 188 _a | 13.2% | 130 _b | 15.3% | 85 _b | 14.8% | 137 _b | 9.3% | 411 _a | 10.7% | 951 | |
| Severe depression (20-27) | 7.6% | 149 _a | 7.6% | 75 _a | 12.6% | 70 _b | 16.1% | 149 _b | 7.2% | 317 _a | 8.6% | 760 | |
| GAD-7 score (0-21) | 5.8 _a | 5.7 | 7.0 _b | 5.7 | 8.3 _c | 6.4 | 8.4 _c | 6.3 | 5.7 _a | 5.7 | 6.3 | 5.9 | < 0.001 |
| GAD-7 severity | | | | | | | | | | | | | |
| None or minimal (0-4) | 51.4% | 1,008 _a | 39.7% | 392 _b | 32.5% | 180 _c | 31.2% | 289 _c | 51.1% | 2,265 _a | 46.7% | 4,134 | |
| Mild anxiety (5-9) | 24.8% | 486 _a | 30.8% | 304 _b | 29.2% | 162 _{a,b} | 28.6% | 265 _{a,b} | 25.2% | 1,118 _a | 26.4% | 2,335 | 0.001 |
| Moderate anxiety (10-14) | 14.3% | 281 _a | 16.9% | 167 _{a,b} | 17.9% | 99 _{a,b} | 20.5% | 190 _b | 14.5% | 641 _a | 15.6% | 1,378 | <0.001 |
| Severe anxiety (15-21) | 9.5% | 187 _{a,c} | 12.6% | 124 _a | 20.4% | 113 _b | 19.7% | 182 _b | 9.1% | 405 _c | 11.4% | 1,011 | |

Values in the same row and subtable not sharing the same subscript are significantly different at P<0.05 in the 2-sided test of equality for column proportions. Cells with no subscript are not included in the test. Tests assume equal variances. Tests are adjusted for all pairwise comparisons within a row of each innermost subtable using the Bonferroni correction.

^aThis category is not used in comparisons because its column proportion is equal to 0 or 1.

^bSum of the following diagnosed conditions: breast cancer, cervical cancer, ovarian cancer, uterine cancer, endometriosis, fibroids, dysmenorrhea, heavy menstrual bleeding, hot flashes, premenstrual dysphoric disorder, and premenstrual syndrome.

BMI=body mass index; FT/PT/SE=full-time/part-time/self-employed; GAD-7=Generalized Anxiety Disorder questionnaire 7-item; HFEM=high-frequency episodic migraine; LFEM=low-frequency episodic migraine; MFEM=moderate-frequency episodic migraine; PHQ-9=Patient Health Questionnaire 9-item; Rx=prescription.

Results

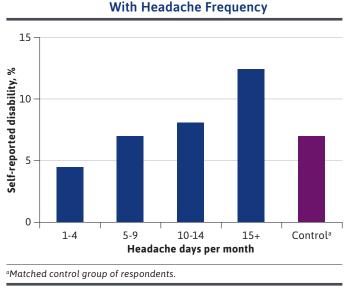
PATIENT DEMOGRAPHICS

Prematch observations in the migraine cohort (N=4,429) indicated that 44.3% experienced LFEM, 22.3% experienced MFEM, 12.5% experienced HFEM, and 20.9% experienced CM. All groups had a similar duration of illness (15.1-15.7 mean years), with the majority (60.8%; 2,660 of 4,487) of respondents having received their diagnosis more than 10 years earlier. Additional information on the self-reported clinical characteristics (eg, symptomology, treatment history) can be found in <u>Supplementary Table 3</u>.

Postmatch analyses, conducted after creation of the propensity-matched groups, revealed several differences between migraine cohorts relative to the matched control group. As shown in Table 1, individuals with CM were more likely to be female (79.1% [vs matched control, 74.3%]), more likely to be White (72.5% [vs matched control, 65.4%]), less likely to report completion of a 4-year degree (35.0% [vs

matched control, 43.9%]), less likely to be employed (54.1% [vs matched control, 64.0%]), and more likely to earn less than \$25k per year (26.8% [vs matched control, 19.2%]) (all P<0.001). Whereas CM was associated with lower socio-economic status relative to the matched control group, individuals with LFEM were characterized by higher socio-economic status than matched control group (completion of a 4-year degree, 51.5%; employed, 71.5%; earning <\$25k per year, 15.7%) (P<0.001). MFEM and HFEM respondents did not consistently differ from the matched control group.

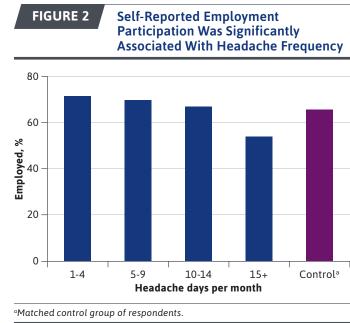
Overall health status was associated with headache frequency, with respondents reporting more frequent headaches that were associated with greater comorbidity burden (mean CCI: LFEM 0.5 ± 1.0 , MFEM 0.6 ± 1.1 , HFEM 0.6 ± 1.2 , CM 0.7 ± 1.2 , matched control 0.6 ± 1.3 ; P<0.001), and more frequent self-reported diagnoses of depression (LFEM 39.1%, MFEM 46.0%, HFEM 48.2%, CM 58.2%, matched control 47.3%; P<0.001) and anxiety (LFEM 35.1%, MFEM 42.5%, HFEM 45.0%, CM 54.0%, matched control 42.3%;



Self-Reported Short- or Long-Term

Disability Was Significantly Associated

FIGURE 1



P<0.001). Headache frequency was also positively associated with the severity of one's depression and anxiety, as measured by the PHQ-9 (LFEM 7.4±7.0, MFEM 8.7±6.9, HFEM 10.1±7.4, CM 11.0±7.6, matched control 7.2±6.9; P<0.001) and GAD-7 (LFEM 5.8±5.7, MFEM 7.0±5.7, HFEM 8.3±6.4, CM 8.4 ± 6.3 , matched control 5.7 ± 5.7 ; P<0.001) scores. Given that more than 70% of our sample was female, we examined diagnoses specific to women's health, including breast, cervical, ovarian, and uterine cancer, endometriosis, fibroids, dysmenorrhea, heavy menstrual bleeding, hot flashes, premenstrual dysphoric disorder, and premenstrual syndrome. Although infrequently reported, these diagnoses were more frequently reported by women with migraine compared with the matched control population, at proportions approximately 2-fold higher among those with CM and HFEM (mean number of diagnosed women's health conditions: LFEM 0.7±1.0, MFEM 0.7±1.1, HFEM 0.8±1.1, CM 0.9±1.2, matched control 0.4±0.8; P<0.001) (Table 1).

Employment Status. As shown in Figure 1, headache frequency was higher in respondents who self-reported short- or long-term disability, which was reported roughly twice as often among those with CM, relative to the matched control group (LFEM 4.4%, MFEM 6.4%, HFEM 7.9%, CM 12.3%, matched control 6.9%; P<0.001). Interestingly, reports of short- or long-term disability among those with LFEM were significantly lower than that observed in the matched control group (P<0.05). This finding was true whether the LFEM cohort was compared with the matched control cohort propensity matched to the overall migraine cohort or with the subset matched control cohort that was propensity matched to LFEM/MFEM cohorts (data available on request). Consistent with this observation, the likelihood of being employed was lower in respondents reporting higher migraine frequency (Figure 2).

MIDAS scores evaluating migraine-related disability were also higher in respondents reporting greater headache frequency (LFEM 11.4 \pm 17.7, MFEM 22.5 \pm 27.5, HFEM 33.1 \pm 37.3, CM 56.5 \pm 58.5; P<0.001) (Table 2), with 49.5% and 64.0% of respondents with HFEM and CM, respectively, being classified as having the most severe level of disability measured by the instrument.

The percentage of individuals who reported either short- or long-term disability represented only a small fraction of respondents (13.9% [614 of 4,429 total migraine cohort]), compared with those who were classified as disabled by MIDAS score (34.2% [1,516 of 4,429 total migraine cohort]) (Figure 3). Whereas 14.5% of respondents with LFEM were characterized as severely disabled according to MIDAS, only 4.4% of this group self-reported having employment status as short- or long-term disability. This disparity was magnified by increased headache frequency (frequency of severe disability by MIDAS vs frequency of self-reported short- or long-term disability observed in the sample: MFEM 36.9% vs 6.4%, HFEM 49.5% vs 7.9%, CM 64.0% vs 12.3%).

| TABLE 2 | Migrai |
|---------|--------|
|---------|--------|

Migraine-Related Disability

| | Migraine frequency (includes HFEM) | | | | | | | | | | | | |
|---|------------------------------------|-----------------------------|-------------------|----------------------------|-------------------|---------------------------------|-------------------|----------------------------------|--------------|---|--------------|----------------------|------------|
| | | LFEM (≤4 days) (n=1,962) | | MFEM (5-9 days) (n=987) | | HFEM (10-14 days) (n=554) | | Chronic (≥15 days) (n=926) | | Matched control (no migraine) (n=4,429) | | Total (N = 8,858) | |
| | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | % or mean | N or SD | P value |
| Disability-related unemploym | ent: empl | oyment | | | | | | | | | | | |
| Employed full-time | 54.9% | 1,078 _a | 50.4% | 497 _{a,b} | 45.5% | 252 _b | 34.1% | 316 _c | 45.9% | 2,034 _{b,d} | 47.2% | 4,177 | |
| Self-employed | 6.7% | 132 _a | 8.2% | 81 _a | 9.2% | 51 _a | 8.4% | 78 _a | 7.2% | 321 _a | 7.5% | 663 | <0.001 |
| Employed part-time | 9.8% | 192 _a | 11.3% | 112 _a | 12.3% | 68 _a | 11.6% | 107 _a | 10.8% | 479 _a | 10.8% | 958 | |
| Homemaker | 7.5% | 147 _a | 9.2% | 91 _{a,b} | 9.6% | 53 _{a,b} | 12.3% | 114 _b | 8.5% | 376 _a | 8.8% | 781 | |
| Retired | 6.7% | 132 _a | 3.5% | 35 _b | 4.3% | 24 _{a,b} | 5.2% | 48 _{a,b} | 6.4% | 282 _a | 5.9% | 521 | |
| Student | 4.13% | 81 _a | 4.4% | 43 _{a,b} | 4.3% | 24 _{a,b} | 6.8% | 63 _b | 6.2% | 273 _{b,c} | 5.5% | 484 | |
| Disability (long-term or short-term) | 4.4% | 86 _a | 6.4% | 63 _{a,b,d} | 7.9% | 44 _{b,c,d} | 12.3% | 114 _c | 6.9% | 307 _d | 6.9% | 614 | |
| Not employed but looking for work | 4.9% | 97 _a | 5.2% | 51 _a | 5.8% | 32 _a | 6.2% | 57 _a | 6.4% | 282 _a | 5.9% | 519 | |
| Not employed and not looking for work | 0.9% | 17 _a | 1.4% | 14 _{a,b} | 1.1% | 6 _{a,b} | 3.1% | 29 _b | 1.7% | 75 _a | 1.6% | 141 | |
| Disability: MIDAS grade | | | | | | | | | | | | | |
| Grade 1: little or no disability (0-5) | 46.5% | 913 _a | 24.6% | 243 _b | 18.6% | 103 _c | 16.2% | 150 _c | NA | NA | 31.8% | 1,409 | |
| Grade 1: mild disability (6-10) | 19.5% | 382 _a | 14.6% | 144 _b | 9.7% | 54 _c | 5.1% | 47 _d | NA | NA | 14.2% | 627 | <0.001 |
| Grade 3: moderate disability (11-20) | 19.5% | 382 _a | 23.9% | 236 _b | 22.2% | 123 _{a,b} | 14.7% | 136 _c | NA | NA | 19.8% | 877 | |
| Grade 4: severe disability (21+) | 14.5% | 285 _a | 36.9% | 364 _b | 49.5% | 274 _c | 64.0% | 593 _d | NA | NA | 34.2% | 1,516 | |
| Disability: MIDAS score | 11.4ª | 17.7 | 22.5 _b | 27.5 | 33.1 _c | 37.3 | 56.5 _d | 58.5 | NA | NA | 26.0 | 38.7 | <0.001 |

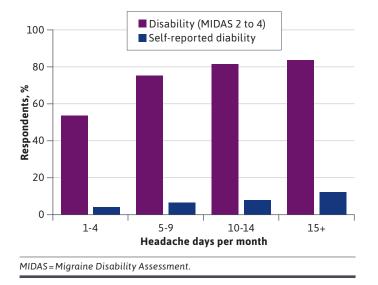
Values in the same row and subtable not sharing the same subscript are significantly different at P<0.05 in the 2-sided test of equality for column proportions. Cells with no subscript are not included in the test. Tests assume equal variances. Tests are adjusted for all pairwise comparisons within a row of each innermost subtable using the Bonferroni correction.

HFEM=high-frequency episodic migraine; LFEM=low-frequency episodic migraine; MFEM=moderate-frequency episodic migraine; MIDAS=Migraine Disability Assessment; NA=not applicable.

Discussion

In this study, we examined the relationship between headache frequency and unemployment, in which the NHWS captured survey data from individuals who self-reported a diagnosis of migraine from a health care professional. We found that those experiencing more than 15 headache days per month were nearly 3 times more likely to report their employment status as disabled as those with fewer than 4 headache days per month. They were also nearly twice as likely to report a disabled status compared with the matched control group sample, which also reported some disability potentially because of other comorbidities. We also examined the relationship between self-reported short- and long-term disability and overall disability as quantified by the MIDAS questionnaire. Our results indicate that both self-reported short- and long-term disability and MIDAS disability tended to be higher in those respondents reporting greater migraine frequency; however, nearly 3 times more people living with migraine are classified as disabled via MIDAS scores as are those self-reporting employmentrelated disability. Overall, the data suggest that people with CM are more likely to report unemployment because of their disease status.

FIGURE 3 The Frequency of Short- or Long-Term Disability and Disability (MIDAS Scores 2-4) Reported Among Adults With Migraine Was Strikingly Disparate, Despite Both Being Positively Associated With Headache Frequency



The disparity between the frequency of MIDAS-assessed severe disability and the self-reported short- and longterm disability suggests that many people living with CM continue to work despite impairments due to perceived or actual barriers to entitlements or accommodations.²⁰ However, people living with migraine might gravitate toward occupations or employers in which accommodations can be made for the challenges faced. Alternatively, people living with migraine may develop coping skills to maintain their employment despite significant disability, or they may have encountered specific legal barriers associated with qualifying for work-related medical disability benefits due to migraine.44,45 In a tight labor market, efforts to reduce headache days among those with migraine can add to the pool of available employees and, therefore, benefit not only the employee but also employers in the United States. Finally, stigma associated with migraine may contribute to the underreporting of short- and long-term disability (and/or failure to seek work-related medical disability benefits).^{45,46} Inability to work has been shown to be the strongest predictor of stigma in migraine⁴⁷; therefore, our findings indicate that exploring the discordance between MIDAS-assessed disability and short- and longterm employment disability might illuminate the impact of this stigma.

An association between poor measures of health with more frequent headaches was observed across study measures that impact disability as defined by self-reported short- and long-term disability and MIDAS. Among adults with migraine, more frequent headaches were associated with higher CCI scores. Anxiety and depression were more prevalent and more severe among women with more frequent headaches (Table 1). Consistent with other epidemiological research, we found higher headache frequency and more coexisting health conditions among cisgender women, including amenorrhea, gynecologic malignancies, and uterine fibroids. The complete extent of this relationship is unclear, but it may warrant synthesis of a specific morbidity index for research in women with migraine. Given the disproportionately higher prevalence of migraine in women,^{48,49} and the role of the female sex hormones in migraine (including menstrual-related migraine, menstrual migraine, and pure menstrual migraine),^{50,51} we explored the relationship between migraine and female-exclusive morbidities.^{52,53} We found higher headache frequency among female respondents who had female-exclusive comorbidities, such as amenorrhea, gynecologic malignancies, and uterine fibroids, compared with the matched control group.

We observed several differences between the HFEM and CM cohorts, namely a higher frequency of self-reported short- and long-term disability and a substantially higher rate of severe disability as measured by MIDAS score. HFEM seemed to represent an inflection point in this sample after which there was a dramatic increase in rate of self-reported disability, potentially suggesting that HFEM should not be included in the expanded definition of CM.54 Prior work has debated the need to distinguish between HFEM and CM respondents, with some studies reporting that 10 headache days per month is substantially burdensome to individuals with migraine, questioning the merits of a threshold of at least 15 headache days.³¹ Others have documented concerns around widening the definition of CM.54 Our findings provide additional evidence that supports early intervention to halt the transition from acute to CM with the goal of preserving function. However, given the cross-sectional nature of the study design, future studies should continue to explore patterns of burden by frequency of headache to discern the potential for the benefits of early intervention.

A wholly novel observation from our analyses was that, relative to the matched control group, individuals with LFEM were less likely to report short- or long-term disability and were more likely to report being currently employed; these results appear to reflect a higher socioeconomic status among those with LFEM than individuals without migraine in the propensity-matched control cohort. Sensitivity analyses confirmed that this difference in disability status between LFEM and the matched control group persisted after accounting for potential confounders, such as comorbidity burden (CCI) and mental health diagnoses (anxiety and depression). Although speculative, this finding is consistent with patients with LFEM developing coping strategies that confer resilience over time, potentially contributing to greater self-management of symptoms and thus a reduction in short- or long-term disability. Future studies should examine the degree to which headache frequency relates to psychological resilience, and how resilience moderates the degree of self-reported severity and symptomology among adults with migraine. Additionally, it would be interesting to determine whether disability scores would be different if patients were on prophylaxis for migraine; therefore, stratifying study data according to the migraine treatment/prophylaxis status should be included in a future study.

LIMITATIONS

There are several limitations to this analysis. First, these data are self-reported and thereby subject to response bias. Although all respondents were asked whether they had been diagnosed with migraine by a health care provider, we could neither ascertain the diagnosis based on symptom frequency and severity nor fully exclude the presence of other overlapping or confounding headache disorders. Similarly, all treatment information was self-reported; therefore, we could not determine the quality of care received or whether respondents adhered to their prescribed regimens. Second, the migraine case definition in this study differs from the ICHD-3 criteria, as the migraine frequency categorization was based on headache days over the previous month, and patients with CM were not required to have at least 8 migraine days. Further, all respondents who reported having had migraines were required to have experienced at least 1 migraine attack in the past month; this feature of the study design may have biased our sample toward more serious cases, especially regarding a lower estimated prevalence of adults with LFEM and a higher estimated prevalence of adults with CM.55 In addition, self-reported short- and long-term employment disability was evaluated in the context of general employment and was not specific to migraine alone; thus, although our results demonstrate that employment disability is higher in respondents with greater headache frequency, we cannot rule out the contribution of comorbidities (other than migraine) that may mediate this relationship. Moreover, we assumed that those who selected short-term or long-term disability self-assessed

that they were unable to work for some period of time because of impairments related to migraine or other existing comorbidities; however, the type, duration, and extent of employment disability are unknown. Although we presume that this measure of employment disability is a reasonable proxy for those who would be eligible for disability benefits, we had no information on patients' benefit status or the etiology of their qualification for benefits in this survey. These considerations aside, there appears to be little incentive to misrepresent one's disability status in a large, anonymized survey like the NHWS, and many reasons to believe that a person living with migraine is the best authority to legitimately comment on their condition and the limitations that it places on their functioning, including the domain of employment. Likewise, any biases relating to self-report are likely systematically distributed across all respondents and unlikely to account for the incremental differences observed here across migraine frequency subgroups. Lastly, respondents were matched on several demographic and health characteristics, and we were able to confirm that the study cohorts were adequately balanced on these potential confounders; nevertheless, cohorts could not be matched on all possible variables on which there may be preexisting baseline differences, and the observed results may thus be influenced, at least in part, by unmeasured variables.

Conclusions

In this study, we found that more frequent migraine headaches were associated with a higher likelihood of unemployment in terms of self-reported short- or longterm disability. This finding can inform policy approaches for reducing the burden associated with impairment and maximizing the employment opportunities of people living with migraine in our society who may benefit from novel preventive treatments. Future research should also evaluate the association between resilience and disease characteristics for employed individuals with migraine.

DISCLOSURES

Robert E Shapiro is a research consultant for Eli Lilly and Lundbeck. Ashley A Martin and Martine C Maculaitis are employees of Cerner Enviza (formerly Kantar Health), which received payment from Lundbeck to conduct the research. Shiven Bhardwaj was an employee of Lundbeck at the time of study and manuscript development. Heather Thomson and Carlton Anderson are employees of Lundbeck. Steven M Kymes is an employee and stockholder of Lundbeck. Financial support for research conducted and manuscript preparation was provided by Lundbeck.

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REFERENCES

1. Altman BM. Definitions, models, classifications, schemes, and applications. In: Albrecht GL, Seelman K, Bury M, eds. Handbook of Disability Studies. SAGE Publications, Inc; 2012:97-122. doi:10.4135/9781412976251.n4

2. World Health Organization. World report on disability. Published 2011. Accessed February 14, 2022. https://www.who.int/teams/ noncommunicable-diseases/sensoryfunctions-disability-and-rehabilitation/ world-report-on-disability

3. World Health Organization. How to Use the ICF: A Practical Manual for Using the International Classification of Functioning, Disability and Health (ICF). Exposure Draft for Comment. World Health Organization; 2013.

4. Centers for Disease Control and Prevention. Disability and health overview. Published 2020. Accessed February 14, 2022. <u>https://www.cdc.gov/ncbddd/</u> <u>disabilityandhealth/disability.html</u>

5. Steiner TJ, Stovner LJ, Jensen R, Uluduz D, Katsarava Z. Migraine remains second among the world's causes of disability, and first among young women: Findings from GBD2019. J Headache Pain. 2020;21(1):137. doi:10.1186/ s10194-020-01208-0 6. Buse DC, Reed ML, Fanning KM, Bostic RC, Lipton RB. Demographics, headache features, and comorbidity profiles in relation to headache frequency in people with migraine: Results of the American Migraine Prevalence and Prevention (AMPP) Study. *Headache*. 2020;60(10):2340-56. <u>doi:10.1111/</u> <u>head.13966</u>

7. Stang P, Korff M von, Galer BS. Reduced labor force participation among primary care patients with headache. J Gen Intern Med. 1998;13:296-302. doi:10.1046/j.1525-1497.1998.00094.x

8. Buse DC, Yugrakh MS, Lee LK, Bell J, Cohen JM, Lipton RB. Burden of illness among people with migraine and ≥4 monthly headache days while using acute and/or preventive prescription medications for migraine. J Manag Care Spec Pharm. 2020;26(10):1334-43. <u>doi:10.18553/</u> jmcp.2020.20100

9. Hu XH, Markson LE, Lipton RB, Stewart WF, Berger ML. Burden of migraine in the United States: Disability and economic costs. Arch Intern Med. 1999;159(8):813-8. <u>doi:10.1001/</u> <u>archinte.159.8.813</u>

10. D'Amico D, Grazzi L, Grignani E, et al; HEADWORK Study Group. HEADWORK Questionnaire: Why do we need a new tool to assess work-related disability in patients with migraine? Headache. 2020;60(2):497-504. <u>doi:10.1111/</u> <u>HEAD.13735</u>

11. Stewart WF, Lipton RB, Kolodner KB, Sawyer J, Lee C, Liberman JN. Validity of the Migraine Disability Assessment (MIDAS) score in comparison to a diarybased measure in a population sample of migraine sufferers. *Pain.* 2000;88(1):41-52. doi:10.1016/S0304-3959(00)00305-5

12. D'Amico D, Tepper SJ, Guastafierro E, et al. Mapping assessments instruments for headache disorders against the ICF biopsychosocial model of health and disability. Int J Environ Res Public Health. 2020;18(1):1-22. doi:10.3390/ijerph18010246 13. Stewart WF, Lipton RB, Dowson AJ, Sawyer J. Development and testing of the Migraine Disability Assessment (MIDAS) Questionnaire to assess headache-related disability. *Neurology*. 2001;56(6 Suppl 1):S20-8. doi:10.1212/wnl.56.suppl 1.S20

14. D'Amico YD, Usai S, Grazzi L, et al. Disability and migraine: MIDAS. J Headache Pain. 2001;2(Suppl 1):s25-7. doi:10.1007/S101940170005

15. Buse DC, Rupnow MFT, Lipton RB. Assessing and managing all aspects of migraine: Migraine attacks, migrainerelated functional impairment, common comorbidities, and quality of life. *Mayo Clin Proc.* 2009;84(5):422-35. <u>doi:10.1016/</u> <u>S0025-6196(11)60561-2</u>

16. Buse DC, Scher AI, Dodick DW, et al. Impact of migraine on the family: Perspectives of people with migraine and their spouse/domestic partner in the CaMEO study. *Mayo Clin Proc.* 2016;91(5):596-611. <u>doi:10.1016/j.</u> <u>mayocp.2016.02.013</u>

17. Lipton RB, Buse DC, Adams AM, Varon SF, Fanning KM, Reed ML. Family impact of migraine: Development of the Impact of Migraine on Partners and Adolescent Children (IMPAC) Scale. *Headache*. 2017;57(4):570–85. <u>doi:10.1111/</u> <u>head.13028</u>

18. Buse DC, Fanning KM, Reed ML, et al. Life with migraine: Effects on relationships, career, and finances from the Chronic Migraine Epidemiology and Outcomes (CaMEO) Study. *Headache*. 2019;59(8):1286-99. <u>doi:10.1111/head.13613</u>

19. Leonardi M, Raggi A. A narrative review on the burden of migraine: When the burden is the impact on people's life. J Headache Pain. 2019;20(1):41. <u>doi:10.1186/</u> <u>s10194-019-0993-0</u>

20. Agosti R. Migraine burden of disease: From the patient's experience to a socioeconomic view. *Headache*. 2018;58(Suppl 1):17-32. <u>doi:10.1111/head.13301</u>

21. Wong LP, Alias H, Bhoo-Pathy N, et al. Impact of migraine on workplace productivity and monetary loss: A study of employees in banking sector in Malaysia. *J Headache Pain*. 2020;21(1):68. <u>doi:10.1186/</u> s10194-020-01144-z 22. Stewart WF, Wood GC, Manack A, Varon SF, Buse DC, Lipton RB. Employment and work impact of chronic migraine and episodic migraine. J Occup Environ Med. 2010;52(1):8-14. <u>doi:10.1097/</u> JOM.0b013e3181c1dc56

23. Gibbs SN, Shah S, Deshpande CG, et al. United States patients' perspective of living with migraine: Country-specific results from the global "My Migraine Voice" Survey. *Headache*. 2020;60(7):1351-64. doi:10.1111/head.13829

24. Gupta S, Richard L, Forsythe A. The humanistic and economic burden associated with increasing body mass index in the EU5. Diabetes Metab Syndr Obes. 2015;8:327-38. doi:10.2147/DMSO.S83696

25. Gupta S, Ryvlin P, Faught E, Tsong W, Kwan P. Understanding the burden of focal epilepsy as a function of seizure frequency in the United States, Europe, and Brazil. *Epilepsia Open*. 2017;2(2):199– 213. <u>doi:10.1002/epi4.12050</u>

26. Gupta S, Kwan P, Faught E, Tsong W, Forsythe A, Ryvlin P. Understanding the burden of idiopathic generalized epilepsy in the United States, Europe, and Brazil: An analysis from the National Health and Wellness Survey. *Epilepsy Behav.* 2016;55:146-56. <u>doi:10.1016/j.</u> <u>yebeh.2015.12.018</u>

27. Gupta S, Isherwood G, Jones K, Van Impe K. Assessing health status in informal schizophrenia caregivers compared with health status in non-caregivers and caregivers of other conditions. BMC Psychiatry. 2015;15:162. doi:10.1186/s12888-015-0547-1

28. Cerner Enviza. National Health and Wellness Survey NHWS. Accessed August 23, 2022. <u>https://www.cernerenviza.com/</u> <u>real-world-data/national-health-and-</u> <u>wellness-survey-nhws</u>

29. Public Policy Committee International Society of Pharmacoepidemiology. Guidelines for good pharmacoepidemiology practice (GPP). Pharmacoepidemiol Drug Saf. 2016;25(1):2-10. <u>doi:10.1002/</u> <u>pds.3891</u> 30. Headache Classification Committee of the International Headache Society (IHS). The International Classification of Headache Disorders, 3rd edition. *Cephalalgia*. 2018;38(1):1-211. <u>doi:10.1177/0333102417738202</u>

31. Torres-Ferrús M, Quintana M, Fernandez-Morales J, Alvarez-Sabin J, Pozo-Rosich P. When does chronic migraine strike? A clinical comparison of migraine according to the headache days suffered per month. *Cephalalgia*. 2017;37(2):104-13. doi:10.1177/0333102416636055

32. Lipton RB, Hutchinson S, Ailani J, et al. Discontinuation of acute prescription medication for migraine: results from the Chronic Migraine Epidemiology and Outcomes (CaMEO) study. *Headache*. 2019;59(10):1762-72. <u>doi:10.1111/head.13642</u>

33. Charlson ME, Pompei P, Ales KL, MacKenzie CR. A new method of classifying prognostic comorbidity in longitudinal studies: Development and validation. J Chronic Dis. 1987;40(5):373-83. doi:10.1016/0021-9681(87)90171-8

34. Burch RC, Buse DC, Lipton RB.
Migraine: Epidemiology, burden, and comorbidity. Neurol Clin. 2019;37(4):631-49. doi:10.1016/j.ncl.2019.06.001

35. Smitherman TA, Rains JC, Penzien DB. Psychiatric comorbidities and migraine chronification. *Curr Pain Headache* Rep. 2009;13(4):326-31. <u>doi:10.1007/</u> <u>s11916-009-0052-7</u>

36. Buse DC, Silberstein SD, Manack AN, Papapetropoulos S, Lipton RB. Psychiatric comorbidities of episodic and chronic migraine. J Neurol. 2013;260(8):1960-9. <u>doi:10.1007/s00415-012-6725-x</u>

37. Caponnetto V, Deodato M, Robotti M, et al; European Headache Federation School of Advanced Studies (EHF-SAS). Comorbidities of primary headache disorders: A literature review with metaanalysis. J Headache Pain. 2021;22(1):71. doi:10.1186/s10194-021-01281-z 38. Ruiz MA, Zamorano E, García-Campayo J, Pardo A, Freire O, Rejas J. Validity of the GAD-7 scale as an outcome measure of disability in patients with generalized anxiety disorders in primary care. J Affect Disord. 2011;128(3):277-86. doi:10.1016/j.jad.2010.07.010

39. Kroenke K, Spitzer RL, Williams JBW. The PHQ-9: Validity of a brief depression severity measure. J Gen Intern Med. 2001;16(9):606-13.

40. Austin PC. An introduction to propensity score methods for reducing the effects of confounding in observational studies. *Multivariate Behav Res*. 2011;46(3):399. <u>doi:10.1080/00273171.2011.</u> 568786

41. Normand ST, Landrum MB, Guadagnoli E, et al. Validating recommendations for coronary angiography following acute myocardial infarction in the elderly: A matched analysis using propensity scores. J Clin Epidemiol. 2001;54(4):387-98. <u>doi:10.1016/</u> <u>s0895-4356(00)00321-8</u>

42. Cohen J. Statistical Power Analysis for the Behavioral Sciences. Routledge; 2013. <u>doi:10.4324/9780203771587</u>

43. Rubin DB. Using propensity scores to help design observational studies: Application to the tobacco litigation. Health Serv Outcomes Res Methodol. 2001;2:169–88. doi:10.1023/A:1020363010465

44. Borkum J, Evans RW. Disability and chronic migraine. *Headache*. 2014;54(4):719-25. <u>doi:10.1111/head.12339</u>

45. Shapiro RE. What will it take to move the needle for headache disorders? An advocacy perspective. *Headache*. 2020;60(9):2059-77. <u>doi:10.1111/head.13913</u>

46. Parikh SK, Young WB. Migraine: Stigma in society. *Curr Pain Headache* Rep. 2019;23(1):8. <u>doi:10.1007/</u> <u>s11916-019-0743-7</u>

47. Young WB, Park JE, Tian IX, Kempner J. The stigma of migraine. PLoS One. 2013;8(1):e54074. <u>doi:10.1371/journal.</u> <u>pone.0054074</u> 48. Lipton RB, Stewart WF, Diamond S, Diamond ML, Reed M. Prevalence and burden of migraine in the United States: Data from the American Migraine Study II. *Headache*. 2001;41(7):646-57. doi:10.1046/j.1526-4610.2001.041007646.x

49. Lipton RB, Bigal ME, Diamond M, Freitag F, Reed ML, Stewart WF; AMPP Advisory Group. Migraine prevalence, disease burden, and the need for preventive therapy. *Neurology*. 2007;68(5):343-9. <u>doi:10.1212/01.wnl.0000252808.97649.21</u>

50. Karlı N, Betül B, Mustafa E, et al; Turkish Headache Prevalence Study Group, Onal AE. Impact of sex hormonal changes on tension-type headache and migraine: A cross-sectional populationbased survey in 2,600 women. J Headache Pain. 2012;13(7):557-65. <u>doi:10.1007/</u> <u>s10194-012-0475-0</u> 51. Dzoljic E, Sipetic S, Vlajinac H, et al. Prevalence of menstrually related migraine and nonmigraine primary headache in female students of Belgrade University. *Headache*. 2002;42(3):185-93. <u>doi:10.1046/J.1526-4610.2002.02050.X</u>

52. Tietjen GE, Conway A, Utley C, Gunning WT, Herial NA. Migraine is associated with menorrhagia and endometriosis. *Headache*. 2006;46(3):422-8. doi:10.1111/j.1526-4610.2006.00290.x

53. Yang MH, Wang PH, Wang SJ, Sun WZ, Oyang YJ, Fuh JL. Women with endometriosis are more likely to suffer from migraines: A population-based study. PLoS One. 2012;7(3):e33941. <u>doi:10.1371/journal.</u> pone.0033941 54. Guglielmetti M, Raggi A, Ornello R, et al. The clinical and public health implications and risks of widening the definition of chronic migraine. *Cephalalgia*. 2020;40(4):407-10. <u>doi:10.1177/0333102419895777</u>

55. Blumenfeld AM, Varon SF, Wilcox TK, et al. Disability, HRQoL and resource use among chronic and episodic migraineurs: Results from the International Burden of Migraine Study (IBMS). *Cephalalgia*. 2011;31(3):301-15. doi:10.1177/0333102410381145